

Tests of Multiple Proportions & Goodness of Fit – Prof. Richard B. Goldstein

CONTINGENCY TABLE

H_0 : Rows and Columns are **Independent**

H_1 : There is some dependency

Observed Values					
	A	B	...	K	Total
1	O_{11}	O_{12}	...	O_{1k}	$R_1 = \sum O_{1j}$
2	O_{21}	O_{22}	...	O_{2k}	$R_2 = \sum O_{2j}$
...
L	O_{L1}	O_{L2}	...	O_{Lk}	$R_L = \sum O_{Lj}$
Total	$C_1 = \sum O_{i1}$	$C_2 = \sum O_{i2}$...	$C_k = \sum O_{ik}$	N

Expected Values

where $N = C_1 + C_2 + \dots + C_k = R_1 + R_2 + \dots + R_L$

Expected Values					
	A	B	...	K	Total
1	E_{11}	E_{12}	...	E_{1k}	$R_1 = \sum E_{1j}$
2	E_{21}	E_{22}	...	E_{2k}	$R_2 = \sum E_{2j}$
...
L	E_{L1}	E_{L2}	...	E_{Lk}	$R_L = \sum E_{Lj}$
Total	$C_1 = \sum E_{i1}$	$C_2 = \sum E_{i2}$...	$C_k = \sum E_{ik}$	N

Expected Values

where $E_{ij} = (R_i C_j) / N$

$$E_{ij} = \frac{R_i \cdot C_j}{N} = \frac{i^{\text{th}} \text{ row sum} \cdot j^{\text{th}} \text{ column sum}}{\text{grand total}}$$

$$\chi^2 = \sum_{i=1}^L \sum_{j=1}^K \frac{(O_{ij} - E_{ij})^2}{E_{ij}} \quad \text{with } (K-1)(L-1) \text{ d.f.}$$

note: a test on K proportions can be a contingency table with L = 2 rows

example

Observed:

	A	B	C	Total
1	45	95	60	200
2	30	48	22	100
3	75	97	128	300
Total	150	240	210	600

Expected:

	A	B	C	Total
1	50	80	70	200
2	25	40	35	100
3	75	120	105	300
Total	150	240	210	600

$$\chi^2 = \frac{(45 - 50)^2}{50} + \frac{(95 - 80)^2}{80} + \frac{(60 - 70)^2}{70} + \dots + \frac{(97 - 120)^2}{120} + \frac{(128 - 105)^2}{105}$$

$$\chi^2 = 0.500 + 2.813 + 1.429 + 1.000 + 1.600 + 4.829 + 0.000 + 4.408 + 5.038 = 21.616$$

with $2 * 2 = 4$ d.f. has a p-value of 0.000239

One can look at terms in the summation for χ^2 to see where the greatest dependency occurred.

MULTIPLE PROPORTIONS

$$H_0 : p_1 = p_2 = \dots = p_k$$

H_1 : not all proportions are equal

$$\chi^2 = \sum \frac{(\text{Obs} - \text{Exp})^2}{\text{Exp}}$$

NOTE: Expected Number should always exceed 5

$$\text{Let } \hat{p}_1 = \frac{r_1}{n_1}, \hat{p}_2 = \frac{r_2}{n_2}, \dots, \hat{p}_k = \frac{r_k}{n_k} \text{ and } \hat{p} = \frac{R}{N} = \frac{r_1 + r_2 + \dots + r_n}{n_1 + n_2 + \dots + n_k}$$

Then,

$$\chi^2 = \frac{(r_1 - n_1\hat{p})^2}{n_1\hat{p}} + \frac{(r_2 - n_2\hat{p})^2}{n_2\hat{p}} + \dots + \frac{(r_k - n_k\hat{p})^2}{n_k\hat{p}} + \frac{(n_1 - r_1 - n_1\hat{q})^2}{n_1\hat{q}} + \dots + \frac{(n_k - r_k - n_k\hat{q})^2}{n_k\hat{q}}$$

$$\chi^2 = \sum_{i=1}^k \frac{(r_i - n_i\hat{p})^2}{n_i\hat{p}} + \sum_{i=1}^k \frac{(n_i - r_i - n_i\hat{q})^2}{n_i\hat{q}} \text{ with } k - 1 \text{ d.f.}$$

example

$$\text{let } \hat{p}_1 = \frac{12}{80} = 0.15, \hat{p}_2 = \frac{19}{100} = 0.19, \hat{p}_3 = \frac{36}{120} = 0.30, \text{ and } \hat{p}_4 = \frac{48}{200} = 0.24$$

$$\text{then } \hat{p} = \frac{12 + 19 + 36 + 48}{80 + 100 + 120 + 200} = \frac{115}{500} = 0.23 \text{ and } \hat{q} = 1 - 0.23 = 0.77$$

$$\chi^2 = \frac{(12 - 18.4)^2}{18.4} + \frac{(19 - 23)^2}{23} + \frac{(36 - 27.6)^2}{27.6} + \frac{(48 - 46)^2}{46} \\ + \frac{(68 - 61.6)^2}{61.6} + \frac{(81 - 77)^2}{77} + \frac{(84 - 92.4)^2}{92.4} + \frac{(152 - 154)^2}{154}$$

$$\chi^2 = 2.226 + 0.696 + 2.557 + 0.087 + 0.665 + 0.208 + 0.764 + 0.026$$

$$\chi^2 = 7.229 \text{ with } k - 1 = 4 - 1 = 3 \text{ d.f. has a p-value of } 0.0649 \text{ (accept } H_0 \text{ if } \alpha = 0.05)$$

Goodness of Fit

One can use $\chi^2 = \sum \frac{(\text{Obs} - \text{Exp})^2}{\text{Exp}}$ in general to test the goodness-of-fit for statistical distributions. The degrees of freedom = # comparison bins - # parameters needed.

Example 1 Single Die is tossed 600 times (distribution should be uniform with $E_i = 100$)

Result	1	2	3	4	5	6
Frequency	97	105	112	96	101	89

Since the sum = 600, the d.f. = 6 - 1 = 5

$$\chi^2 = \frac{(97-100)^2}{100} + \frac{(105-100)^2}{100} + \dots + \frac{(89-100)^2}{100} = 0.09 + 0.25 + 1.44 + 0.16 + 0.01 + 1.21$$

$$\chi^2 = 3.16 \text{ which as a p-value of } 0.675 \text{ for } 5 \text{ d.f. (that is, this is an honest/fair die)}$$

Example 2 Rutherford and Geiger's Experiment with Radioactive Particles

Let X = alpha particles per 4 second interval. This is similar to the results obtained by E. Rutherford and H. Geiger in one of their classical experiments in 1910. The distribution is believed to approximate a Poisson distribution with the mean of $\lambda = 2.064$ particles every 4 seconds.

Number of Particles	Observed	Expected
0	111	126.2
1	272	261.2
2	292	270.3
3	173	186.5
4	89	96.5
5	45	40.0
6	12	13.8
7	6	5.5
Total	1000	1000

$$\lambda = 0(0.111) + 1(.272) + 2(.292) + 3(.173) + 4(.089) + 5(.045) + 6(.012) + 7(.006) = 2.07$$

$$E_k = 1000 * \frac{(2.07)^k e^{-2.07}}{k!} \text{ (note: } E_7 \text{ was chosen so that the sum was still 1000)}$$

$$\chi^2 = 1.831 + 0.447 + 1.742 + 0.977 + 0.583 + 0.625 + 0.235 + 0.045 = 6.485$$

which with 8 - 2 = 6 d.f. has a p-value of 0.371 (a good fit for Poisson)

note: lose 1 for the sum and 1 for using the estimated parameter, λ , for the expected value

COMMENT: There are better tests for determining if a certain distribution is present. A non-parametric test is Kolmogorov & Smirnov's Test. But even better tests are available for certain distributions by Anderson & Darling or Shapiro & Wilk. For the normal distribution use tests by Geary and D'Agostino.